



# **A Handbook of Statistical Analyses Using R**

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# Analysis of Variance: Weight Gain, Foster Feeding in Rats, Water Hardness and Male Egyptian Skulls

## 4.1 Introduction

## 4.2 Analysis of Variance

## 4.3 Analysis Using R

### 4.3.1 Weight Gain in Rats

Before applying analysis of variance to the data in Table ?? we should try to summarise the main features of the data by calculating means and standard deviations and by producing some hopefully informative graphs. The data is available in the *data.frame* `weightgain`. The following R code produces the required summary statistics

```
R> data("weightgain", package = "HSAUR")
R> tapply(weightgain$weightgain,
+         list(weightgain$source, weightgain$type), mean)
               High Low
Beef      100.0 79.2
Cereal   85.9 83.9

R> tapply(weightgain$weightgain,
+         list(weightgain$source, weightgain$type), sd)
               High      Low
Beef  15.13642 13.88684
Cereal 15.02184 15.70881
```

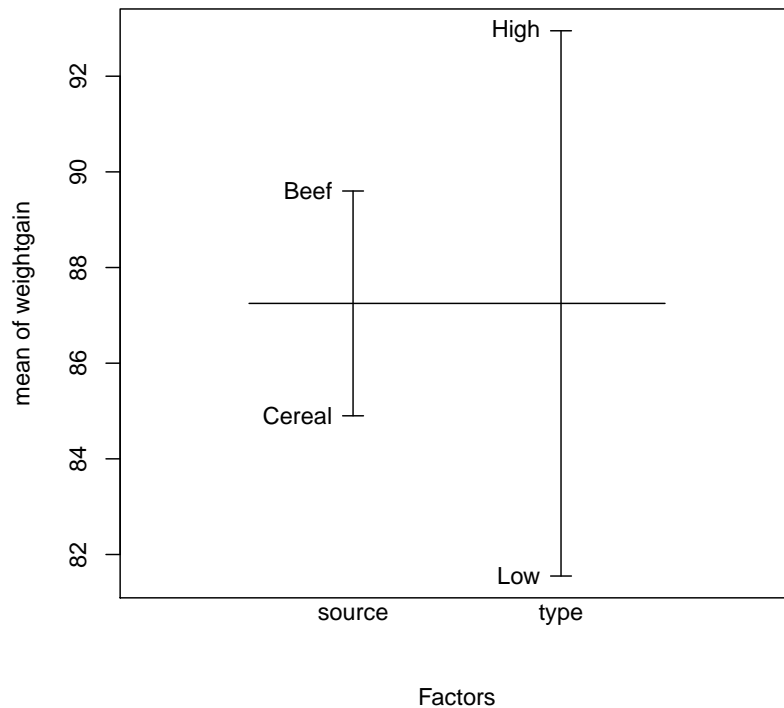
To apply analysis of variance to the data we can use the `aov` function in R and then the `summary` method to give us the usual analysis of variance table. The model *formula* specifies a two-way layout with interaction terms, where the first factor is `source`, and the second factor is `type`.

```
R> wg_aov <- aov(weightgain ~ source * type, data = weightgain)
```

The estimates of the intercept and the main and interaction effects can be extracted from the model fit by

```
R> coef(wg_aov)
               (Intercept)               sourceCereal               typeLow
               100.0                -14.1                -20.8
sourceCereal:typeLow
               18.8
```

```
R> plot.design(weightgain)
```



**Figure 4.1** Plot of mean weight gain for each level of the two factors.

Note that the model was fitted with the restrictions  $\gamma_1 = 0$  (corresponding to Beef) and  $\beta_1 = 0$  (corresponding to High) because treatment contrasts were used as default as can be seen from

```
R> options("contrasts")
```

```
$contrasts
      unordered      ordered
"contr.treatment"  "contr.poly"
```

Thus, the coefficient for **source** of  $-14.1$  can be interpreted as an estimate of the difference  $\gamma_2 - \gamma_1$ . Alternatively, we can use the restriction  $\sum_i \gamma_i = 0$  by

```
R> coef(aov(weightgain ~ source + type + source:type,
+          data = weightgain, contrasts = list(source = contr.sum)))
```

(Intercept)	source1	typeLow
92.95	7.05	-11.40

---

```
R> summary(wg_aov)
```

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
source	1	221	220.9	0.988	0.3269
type	1	1300	1299.6	5.812	0.0211 *
source:type	1	884	883.6	3.952	0.0545 .
Residuals	36	8049	223.6		

```
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

---

**Figure 4.2** R output of the ANOVA fit for the `weightgain` data.

```
source1:typeLow
-9.40
```

#### 4.3.2 Foster Feeding of Rats of Different Genotype

As in the previous subsection we will begin the analysis of the foster feeding data in Table ?? with a plot of the mean litter weight for the different genotypes of mother and litter (see Figure 4.4). The data are in the *data.frame* `foster`

```
R> data("foster", package = "HSAUR")
```

We can derive the two analyses of variance tables for the foster feeding example by applying the R code

```
R> summary(aov(weight ~ litgen * motgen, data = foster))
```

to give

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
litgen	3	60.2	20.05	0.370	0.77522
motgen	3	775.1	258.36	4.763	0.00574 **
litgen:motgen	9	824.1	91.56	1.688	0.12005
Residuals	45	2440.8	54.24		

```
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

and then the code

```
R> summary(aov(weight ~ motgen * litgen, data = foster))
```

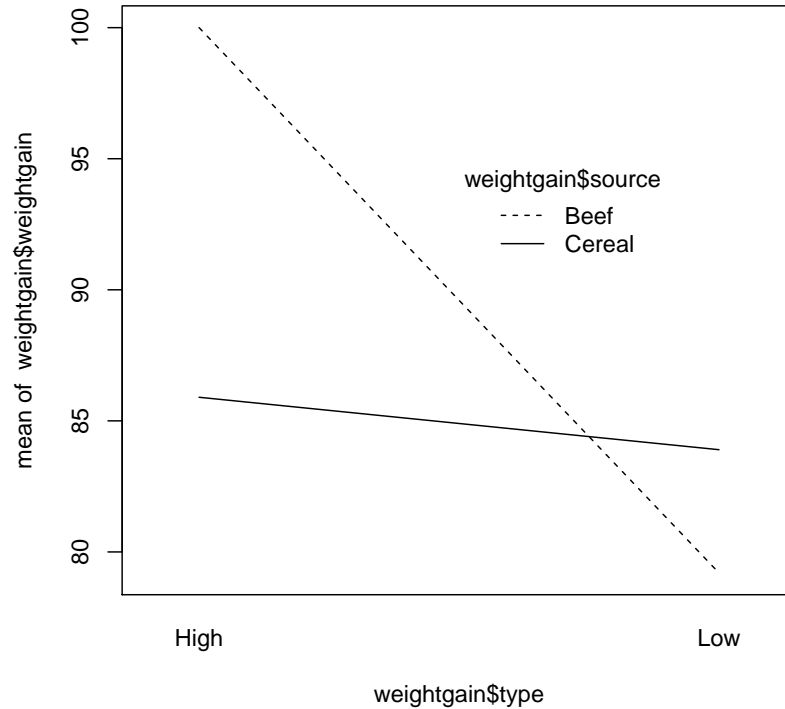
to give

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
motgen	3	771.6	257.20	4.742	0.00587 **
litgen	3	63.6	21.21	0.391	0.76000
motgen:litgen	9	824.1	91.56	1.688	0.12005
Residuals	45	2440.8	54.24		

```
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

There are (small) differences in the sum of squares for the two main effects and, consequently, in the associated  $F$ -tests and  $p$ -values. This would not be true if in the previous example in Subsection 4.3.1 we had used the code

```
R> interaction.plot(weightgain$type, weightgain$source,
+                   weightgain$weightgain)
```



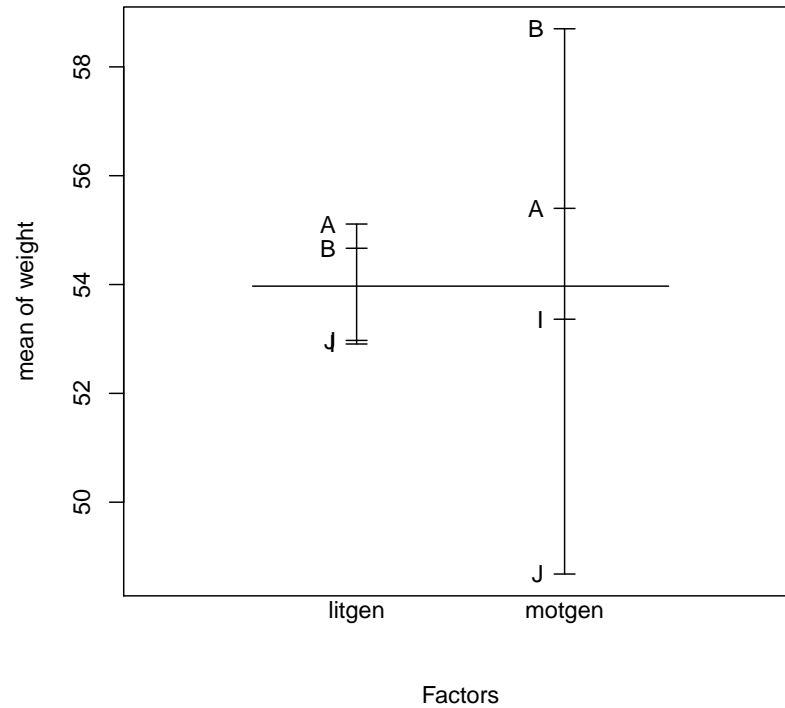
**Figure 4.3** Interaction plot of type  $\times$  source.

```
R> summary(aov(weightgain ~ type * source, data = weightgain))
```

instead of the code which produced Figure 4.2 (readers should confirm that this is the case).

We can investigate the effect of genotype B on litter weight in more detail by the use of *multiple comparison procedures* (see [Everitt, 1996](#)). Such procedures allow a comparison of all pairs of levels of a factor whilst maintaining the nominal significance level at its selected value and producing adjusted confidence intervals for mean differences. One such procedure is called *Tukey honest significant differences* suggested by [Tukey \(1953\)](#), see [Hochberg and Tamhane \(1987\)](#) also. Here, we are interested in simultaneous confidence intervals for the weight differences between all four genotypes of the mother. First, an ANOVA model is fitted

```
R> plot.design(foster)
```



**Figure 4.4** Plot of mean litter weight for each level of the two factors for the `foster` data.

```
R> foster_aov <- aov(weight ~ litgen * motgen, data = foster)
which serves as the basis of the multiple comparisons, here with allpair differ-
ences by
```

```
R> foster_hsd <- TukeyHSD(foster_aov, "motgen")
```

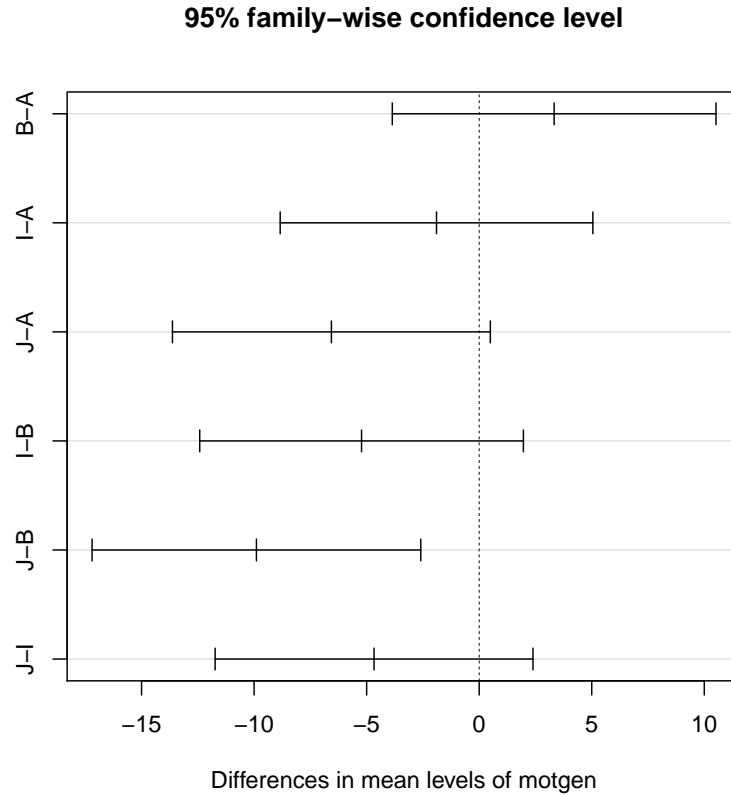
```
R> foster_hsd
```

```
Tukey multiple comparisons of means
95% family-wise confidence level
```

```
Fit: aov(formula = weight ~ litgen * motgen, data = foster)
```

```
$motgen
      diff      lwr      upr    p adj
B-A 3.330369 -3.859729 10.5204672 0.6078581
```

```
R> plot(foster_hsd)
```



**Figure 4.5** Graphical presentation of multiple comparison results for the `foster` feeding data.

```
I-A -1.895574 -8.841869 5.0507207 0.8853702
J-A -6.566168 -13.627285 0.4949498 0.0767540
I-B -5.225943 -12.416041 1.9641552 0.2266493
J-B -9.896537 -17.197624 -2.5954489 0.0040509
J-I -4.670593 -11.731711 2.3905240 0.3035490
```

A convenient `plot` method exists for this object and we can get a graphical representation of the multiple confidence intervals as shown in Figure 4.5. It appears that there is only evidence for a difference in the B and J genotypes.

#### 4.3.3 Water Hardness and Mortality

The water hardness and mortality data for 61 large towns in England and Wales (see Table 2.3) was analysed in Chapter 2 and here we will extend the



analysis by an assessment of the differences of both hardness and mortality in the North or South. The hypothesis that the two-dimensional mean-vector of water hardness and mortality is the same for cities in the North and the South can be tested by *Hotelling-Lawley* test in a multivariate analysis of variance framework. The R function `manova` can be used to fit such a model and the corresponding `summary` method performs the test specified by the `test` argument

```
R> data("water", package = "HSAUR")
R> summary(manova(cbind(hardness, mortality) ~ location,
+ data = water), test = "Hotelling-Lawley")

              Df Hotelling-Lawley approx F num Df den Df      Pr(>F)
location    1          0.90021    26.106      2    58 8.217e-09
Residuals  59

location    ***
Residuals
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

The `cbind` statement in the left hand side of the formula indicates that a *multivariate* response variable is to be modelled. The  $p$ -value associated with the *Hotelling-Lawley* statistic is very small and there is strong evidence that the mean vectors of the two variables are not the same in the two regions. Looking at the sample means

```
R> tapply(water$hardness, water$location, mean)

      North      South 
30.40000  69.76923 

R> tapply(water$mortality, water$location, mean)

      North      South 
1633.600  1376.808
```

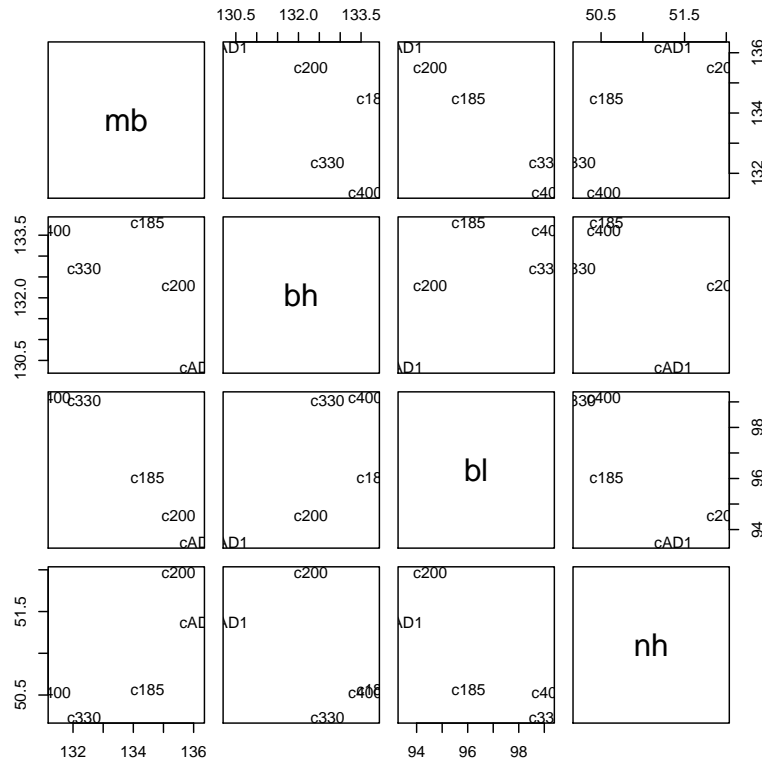
we see large differences in the two regions both in water hardness and mortality, where low mortality is associated with hard water in the South and high mortality with soft water in the North (see Figure ?? also).

#### 4.3.4 Male Egyptian Skulls

We can begin by looking at a table of mean values for the four measurements within each of the five epochs. The measurements are available in the *data.frame* `skulls` and we can compute the means over all epochs by

```
R> data("skulls", package = "HSAUR")
R> means <- aggregate(skulls[,c("mb", "bh", "bl", "nh")],
+ list(epoch = skulls$epoch), mean)
R> means
```

```
R> pairs(means[,-1],
+       panel = function(x, y) {
+         text(x, y, abbreviate(levels(skulls$epoch)))
+       })
```



**Figure 4.6** Scatterplot matrix of epoch means for Egyptian `skulls` data.

	<i>epoch</i>	<i>mb</i>	<i>bh</i>	<i>bl</i>	<i>nh</i>
1	<i>c4000BC</i>	131.3667	133.6000	99.16667	50.53333
2	<i>c3300BC</i>	132.3667	132.7000	99.06667	50.23333
3	<i>c1850BC</i>	134.4667	133.8000	96.03333	50.56667
4	<i>c200BC</i>	135.5000	132.3000	94.53333	51.96667
5	<i>cAD150</i>	136.1667	130.3333	93.50000	51.36667

It may also be useful to look at these means graphically and this could be done in a variety of ways. Here we construct a scatterplot matrix of the means using the code attached to Figure 4.6.

There appear to be quite large differences between the epoch means, at least on some of the four measurements. We can now test for a difference more formally by using MANOVA with the following R code to apply each of the four possible test criteria mentioned earlier;

```
R> skulls_manova <- manova(cbind(mb, bh, bl, nh) ~ epoch,
+                           data = skulls)
R> summary(skulls_manova, test = "Pillai")

              Df  Pillai approx F num Df den Df    Pr(>F)
epoch          4 0.35331    3.512    16   580 4.675e-06 ***
Residuals 145
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

R> summary(skulls_manova, test = "Wilks")

              Df  Wilks approx F num Df den Df    Pr(>F)
epoch          4 0.66359    3.9009    16 434.45 7.01e-07 ***
Residuals 145
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

R> summary(skulls_manova, test = "Hotelling-Lawley")

              Df Hotelling-Lawley approx F num Df den Df
epoch          4          0.48182    4.231    16   562
Residuals 145
              Pr(>F)
epoch      8.278e-08 ***
Residuals
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

R> summary(skulls_manova, test = "Roy")

              Df   Roy approx F num Df den Df    Pr(>F)
epoch          4 0.4251    15.41     4   145 1.588e-10 ***
Residuals 145
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

The  $p$ -value associated with each four test criteria is very small and there is strong evidence that the skull measurements differ between the five epochs. We might now move on to investigate which epochs differ and on which variables. We can look at the univariate  $F$ -tests for each of the four variables by using the code

```
R> summary.aov(skulls_manova)

Response mb :
              Df  Sum Sq Mean Sq F value    Pr(>F)
epoch          4  502.83  125.707   5.9546 0.0001826 ***
Residuals    145 3061.07   21.111
---

```

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

Response bh :

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
epoch	4	229.9	57.477	2.4474	0.04897 *
Residuals	145	3405.3	23.485		

---

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

Response bl :

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
epoch	4	803.3	200.823	8.3057	4.636e-06 ***
Residuals	145	3506.0	24.179		

---

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

Response nh :

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
epoch	4	61.2	15.300	1.507	0.2032
Residuals	145	1472.1	10.153		

We see that the results for the maximum breadths (mb) and basialveolar length (bl) are highly significant, with those for the other two variables, in particular for nasal heights (nh), suggesting little evidence of a difference. To look at the pairwise multivariate tests (any of the four test criteria are equivalent in the case of a one-way layout with two levels only) we can use the `summary` method and `manova` function as follows:

```
R> summary(manova(cbind(mb, bh, bl, nh) ~ epoch, data = skulls,
+ subset = epoch %in% c("c4000BC", "c3300BC")))
```

	Df	Pillai	approx	F num	Df den	Df	Pr(>F)
epoch	1	0.027674	0.39135		4	55	0.8139
Residuals	58						

```
R> summary(manova(cbind(mb, bh, bl, nh) ~ epoch, data = skulls,
+ subset = epoch %in% c("c4000BC", "c1850BC")))
```

	Df	Pillai	approx	F num	Df den	Df	Pr(>F)
epoch	1	0.18757	3.1744		4	55	0.02035 *
Residuals	58						

---

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

```
R> summary(manova(cbind(mb, bh, bl, nh) ~ epoch, data = skulls,
+ subset = epoch %in% c("c4000BC", "c200BC")))
```

	Df	Pillai	approx	F num	Df den	Df	Pr(>F)
epoch	1	0.30297	5.9766		4	55	0.0004564 ***
Residuals	58						

---

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

```
R> summary(manova(cbind(mb, bh, bl, nh) ~ epoch, data = skulls,
+                   subset = epoch %in% c("c4000BC", "cAD150")))

              Df  Pillai approx F num Df den Df    Pr(>F)
epoch          1 0.36182    7.7956      4    55 4.736e-05 ***
Residuals 58
---
```

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

To keep the overall significance level for the set of all pairwise multivariate tests under some control (and still maintain a reasonable power), [Stevens \(2001\)](#) recommends setting the nominal level  $\alpha = 0.15$  and carrying out each test at the  $\alpha/m$  level where  $m$  is the number of tests performed. The results of the four pairwise tests suggest that as the epochs become further separated in time the four skull measurements become increasingly distinct.



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## Bibliography

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- Everitt, B. S. (1996), *Making Sense of Statistics in Psychology: A Second-Level Course*, Oxford, UK: Oxford University Press.
- Hochberg, Y. and Tamhane, A. C. (1987), *Multiple Comparison Procedures*, New York, USA: John Wiley & Sons.
- Stevens, J. (2001), *Applied Multivariate Statistics for the Social Sciences*, Mahwah, New Jersey, USA: Lawrence Erlbaum, 4th edition.
- Tukey, J. W. (1953), “The problem of multiple comparisons (unpublished manuscript),” in *The Collected Works of John W. Tukey VIII. Multiple Comparisons: 1948-1983*, New York, USA: Chapman & Hall.